

Complex Survey Variance and Design Effects in R using the Rstan and Survey packages

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Population Inference from Complex Survey Samples

- ▶ **Goal:** perform **inference** about a finite **population** generated from an unknown **model**, P_{θ_0} .
- ▶ **Data:** from under a **complex sampling design** distribution, P_{ν}
 - ▶ **Probabilities** of inclusion π_i are often **associated with** the **variable** of interest (purposefully)
 - ▶ Sampling designs are “**informative**”: the **balance** of information in the **sample** \neq **balance** in the **population**.
- ▶ **Biased Estimation:** estimate P_{θ_0} **without** accounting for P_{ν} .
 - ▶ Use **inverse probability** weights $w_i = 1/\pi_i$ to **mitigate** bias.
- ▶ **Incorrect Uncertainty Quantification:**
 - ▶ Failure to account for dependence induced by P_{ν} leads to standard errors and confidence intervals that are the **wrong size**.

Variance Estimation

- ▶ The de-facto approach:
 - ▶ approximate sampling **independence** of the primary sampling units (Heeringa et al. 2010).
 - ▶ within-cluster dependence treated as **nuisance**
- ▶ Two common methods:
 - ▶ Taylor **linearization** and **replication** based methods.
 - ▶ A **variety** of implementations are available (Binder 1996, Rao et al. 1992).

Taylor Linearization

Let y_{ij} , X_{ij} , and w_{ij} be the observed data for individual i in cluster j of the sample. Assume the parameter θ is a vector of dimension d with population model value θ_0 .

1. **Approximate** an estimate $\hat{\theta}$, or a 'residual' ($\hat{\theta} - \theta_0$), as a **weighted sum**: $\hat{\theta} \approx \sum_{i,j} w_{ij} z_{ij}(\theta)$ where z_{ij} is a function evaluated at the **current values** of y_{ij} , X_{ij} , and $\hat{\theta}$.

2. **Compute** the weighted components for **each cluster** (e.g., primary sampling units (PSUs)): $\hat{\theta}_j = \sum_i w_{ij} z_{ij}(\theta)$.

3. **Compute** the variance **between** clusters:

$$\widehat{Var}(\hat{\theta}) = \frac{1}{J-d} \sum_{j=1}^J (\hat{\theta} - \hat{\theta}_j)(\hat{\theta} - \hat{\theta}_j)^T$$

4. For stratified designs, compute $\hat{\theta}_s$ and $\widehat{Var}(\hat{\theta}_s)$ **within** strata and sum

$$\widehat{Var}(\hat{\theta}) = \sum_s \widehat{Var}(\hat{\theta}_s).$$

Replication

Let y_{ij} , X_{ij} , and w_{ij} be the observed data for individual i in cluster j of the sample. Assume the parameter θ is a vector of dimension d with population model value θ_0 .

1. Through **randomization** (bootstrap), **leave-one-out** (jackknife), or **orthogonal contrasts** (balanced repeated replicates), create a **set of K replicate weights** $(w_i)_k$ for all $i \in S$ and for every $k = 1, \dots, K$.
2. Each set of weights has a **modified value** (usually 0) for a subset of clusters, and typically has a **weight adjustment** to the other clusters to compensate: $\sum_{i \in S} (w_i)_k = \sum_{i \in S} w_i$ for every k .
3. Estimate $\hat{\theta}_k$ for **each** replicate $k \in 1, \dots, K$.
4. Compute the variance **between** replicates:
$$\widehat{Var}(\hat{\theta}) = \frac{1}{K-d} \sum_{k=1}^K (\hat{\theta} - \hat{\theta}_k)(\hat{\theta} - \hat{\theta}_k)^T.$$
5. For stratified designs, generate replicates such that **each** strata is represented in **every** replicate.

Challenges

There are **two notable trade-offs** associated with these methods:

- ▶ Taylor linearization: value $\hat{\theta}$ computed **once** then used in a plug in for $z_i(\theta)$.
 - ▶ Replication methods: estimate $\hat{\theta}_k$ **computed K times**.
 - ▶ Sizable differences in **computational effort**
- ▶ Replication methods: **no derivatives** are needed.
 - ▶ Taylor linearization: requires the calculation of a **gradient** to derive the **analytical form** of the first order approximation $z_i(\theta)$.
 - ▶ This poses significant **analytical challenges** for all but the simplest models.

Some Improvements

- ▶ **Abstraction of Derivatives** (less analytic work for Taylor Linearization)
 - ▶ Recent advances in **algorithmic differentiation** (Margossian 2018), allows us to specify the model as a log density but only treat the gradient in the abstract **without** specifying it analytically.
 - ▶ Implemented in **Stan** and **Rstan** (Carpenter 2015, Stan Development Team 2016)
- ▶ **Hybrid Approach** or Taylor Linearization for replicate designs (less computation for Replication approaches)
 - ▶ Survey package (Lumley 2016) to calculate replication **variance of gradient**
 - ▶ Use plug in for θ , only estimate **once**

Example: Design Effect for Survey-Weighted Bayes

- ▶ Williams & Savitsky (2018): <https://arxiv.org/abs/1807.11796>
- ▶ Pseudo posterior \propto Pseudo Likelihood \times Prior

$$\pi^\pi(\boldsymbol{\lambda}|\mathbf{y}, \tilde{\mathbf{w}}) \propto \left[\prod_{i=1}^n \pi(y_i|\boldsymbol{\lambda})^{\tilde{w}_i} \right] \pi(\boldsymbol{\lambda})$$

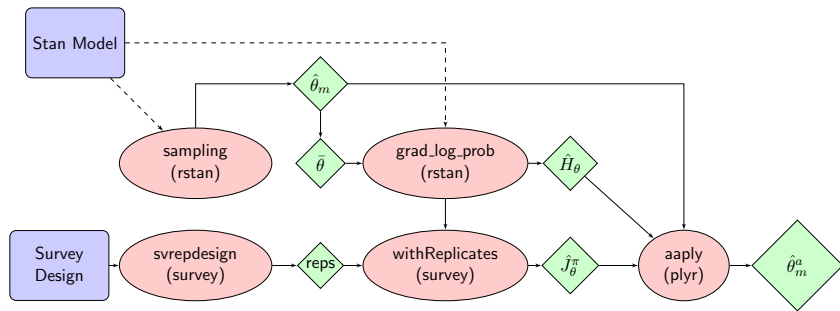
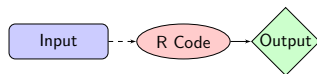
▶ Variances Differ:

- ▶ Weighted MLE: $H_{\theta_0}^{-1} J_{\theta_0}^\pi H_{\theta_0}^{-1}$ (Robust)
- ▶ Weighted Posterior: $H_{\theta_0}^{-1}$ (Model-Based)

▶ Adjust for Design Effect: $R_2^{-1} R_1$

- ▶ $\hat{\theta}_m \equiv$ sample pseudo posterior for $m = 1, \dots, M$ draws with mean $\bar{\theta}$
- ▶ $\hat{\theta}_m^a = (\hat{\theta}_m - \bar{\theta}) R_2^{-1} R_1 + \bar{\theta}$
- ▶ where $R_1' R_1 = H_{\theta_0}^{-1} J_{\theta_0}^\pi H_{\theta_0}^{-1}$
- ▶ $R_2' R_2 = H_{\theta_0}^{-1}$

R Code Schematic



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