#### Discussion of Talk by Jonathan Wright

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The views expressed here are those of the author and not those of the U.S. Census Bureau.

#### Comment on 2 topics covered by Jonathan Wright:

- Comparing MSEs of X-11 and canonical ARIMA (SEATS) seasonal adjustments
- Residual seasonality in NIPA data

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- Do the advantages of indirect SA (of GDP) offset the disadvantage of possible residual seasonality that could potentially be avoided by direct SA?

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Instead of an ARMAX form as above, why not use

$$y_t = \alpha + \beta_1 D_{1t} + \beta_2 D_{2t} + \beta_3 D_{3t} + z_t \qquad (1 - \rho B) z_t = \varepsilon_t$$

# 1. Comparing MSEs of X-11 and canonical ARIMA (SEATS) seasonal adjustments

#### JW Conclusions from Monte Carlo Simulation

- X-13 automatic filter selection tends to select too short seasonal MAs
- Conclusions consistent with other literature
- Model-based SA does better than X-11
- X-11 can get close in some cases
  - But not if  $\theta_{12}$  is close to zero

Compare and contrast results and conclusions with those of

- Chu, Tiao, and Bell (2012) for infinite symmetric filters
- Bell, Chu, and Tiao (2012) for infinite concurrent filters and finite filters

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  $t = 1, \ldots, n$ 

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$$\Rightarrow MSE(\tilde{S}_t) \equiv E[(S_t - \tilde{S}_t)^2] = E[(S_t - \hat{S}_t)^2] + E[(\hat{S}_t - \tilde{S}_t)^2] = g_1 + g_2$$

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$$g_1 = E[(S_t - \hat{S}_t)^2]$$
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where

- $g_1 = E[(S_t \hat{S}_t)^2]$  MSE of optimal predictor
- $g_3 = E[(\hat{S}_t \tilde{S}_t)^2]$  MS difference of  $\tilde{S}_t$  from optimal predictor  $\hat{S}_t$ .

#### Notes on components of seasonal adjustment MSE

- For  $\tilde{S}_t$  a model-based predictor of  $S_t$ ,  $g_3$  reflects the effects of
  - parameter estimation error
  - model selection error (which changes the canonical decomposition)
- For  $\tilde{S}_t$  from X-11 adjustment,  $g_3$  reflects the effects of
  - model selection error and parameter estimation error (affects only forecast extension minor)
  - difference between X-11 filter and optimal model-based filter
    - find which X-11 filter choice minimizes this error

#### Notes on components of seasonal adjustment MSE

- Recall that seasonal adjustment MSE is  $E[(S_t \tilde{S}_t)^2] = g_1 + g_3$ .
  - For any given model,  $g_1$  is the same for any predictor  $\tilde{S}_t$ , while  $g_3$  varies with  $\tilde{S}_t$
- JW estimates  $g_3$  by simulation
  - reports results on  $\sqrt{g_3}$  and ignores  $g_1$
- We ignore  $g_3$  for model-based adjustment, and for X-11 adjustment our  $g_3$  ignores model selection error and parameter estimation error.
  - report MSEs and % differences in MSE between X-11 and optimal model-based adjustment:

$$\mathsf{MSE}~\%~\mathsf{difference}~=100\times\left(\frac{g_1+g_3}{g_1}-1\right)=100\times\left(\frac{g_3}{g_1}\right)$$

• scaling  $g_3$  by  $100/g_1$  aids interpretation of the results

## Other differences between the two approaches to comparisons

Jonathan Wright	Bell, Chu, & Tiao	
reports RMSEs	reports MSEs	
MSEs calculated by simulation	MSEs calculated by analytical formulas	
averages results over $t=1,\ldots,n$	separate results for $t = n/2$ and $t = n$	
??	use full forecast extension for X-11	
$\mathit{n}=120~(10~{ m years})$	results for 8, 12, 16, 20, 40, $\infty$ years	
include X-11 stable seasonal filter	include X-11 3 $ imes$ 15 seasonal MA	

## Comparing MSEs for X-11 and Model-based Filters Canonical decomposition of the airline model with $\theta_1 = .5$ infinite filter results

	$\theta_{12}$			
	.2	.5	.8	.9
Best X-11 seasonal MA	3 × 1	3 × 5	3 × 15	3 × 15
MSE % increase for X-11 symmetric filter concurrent filter	14% 3%	6% 1%	10% 3%	33% 9%

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- Lone exception where best X-11 filter does poorly: seasonal adjustment in the middle of a very long series when θ<sub>12</sub> is large (.9).
- Other X-11 filters with a seasonal MA close to the best choice (for example,  $3 \times 3$  when  $\theta_{12} = .5$ ) have only slightly larger MSEs. X-11 filters far from the best can have larger MSE increases.

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